Statistics for Applications

Chapter 8: Bayesian Statistics

The Bayesian approach (1)

- ▶ So far, we have studied the frequentist approach of statistics.
- The frequentist approach:
 - Observe data
 - These data were generated randomly (by Nature, by measurements, by designing a survey, etc...)
 - We made assumptions on the generating process (e.g., i.i.d., Gaussian data, smooth density, linear regression function, etc...)
 - The generating process was associated to some object of interest (e.g., a parameter, a density, etc...)
 - This object was unknown but fixed and we wanted to find it: we either estimated it or tested a hypothesis about this object, etc...

The Bayesian approach (2)

Now, we still observe data, assumed to be randomly generated by some process. Under some assumptions (e.g., parametric distribution), this process is associated with some fixed object.

- We have a **prior belief** about it.
- Using the data, we want to update that belief and transform it into a **posterior belief**.

The Bayesian approach (3)

Example

- Let p be the proportion of woman in the population.
- ▶ Sample *n* people randomly with replacement in the population and denote by *X*₁,...,*X*_n their gender (1 for woman, 0 otherwise).
- In the frequentist approach, we estimated p (using the MLE), we constructed some confidence interval for p, we did hypothesis testing (e.g., H₀ : p = .5 v.s. H₁ : p ≠ .5).
- Before analyzing the data, we may believe that p is likely to be close to 1/2.
- The Bayesian approach is a tool to:
 - 1. include mathematically our prior belief in statistical procedures.
 - 2. update our prior belief using the data.

The Bayesian approach (4)

Example (continued)

- Our prior belief about p can be quantified:
- ► E.g., we are 90% sure that p is between .4 and .6, 95% that it is between .3 and .8, etc...
- Hence, we can model our prior belief using a distribution for p, as if p was random.
- In reality, the true parameter is not random ! However, the Bayesian approach is a way of modeling our belief about the parameter by doing as if it was random.

• E.g.,
$$p \sim \mathcal{B}(a, a)$$
 (Beta distribution) for some $a > 0$.

• This distribution is called the *prior distribution*.

The Bayesian approach (5)

Example (continued)

- In our statistical experiment, X₁,..., X_n are assumed to be i.i.d. Bernoulli r.v. with parameter p conditionally on p.
- ► After observing the available sample X₁,..., X_n, we can update our belief about p by taking its distribution conditionally on the data.
- The distribution of p conditionally on the data is called the posterior distribution.
- Here, the posterior distribution is

$$\mathcal{B}\left(a+\sum_{i=1}^{n}X_{i},a+n-\sum_{i=1}^{n}X_{i}\right).$$

The Bayes rule and the posterior distribution (1)

- Consider a probability distribution on a parameter space Θ with some pdf π(·): the prior distribution.
- Let X_1, \ldots, X_n be a sample of n random variables.
- ► Denote by $p_n(\cdot|\theta)$ the joint pdf of X_1, \ldots, X_n conditionally on θ , where $\theta \sim \pi$.
- ► Usually, one assumes that X₁,..., X_n are i.i.d. conditionally on θ.
- ► The conditional distribution of θ given X₁,..., X_n is called the *posterior distribution*. Denote by π(·|X₁,...,X_n) its pdf.

The Bayes rule and the posterior distribution (2)

Bayes' formula states that:

$$\pi(\theta|X_1,\ldots,X_n) \propto \pi(\theta)p_n(X_1,\ldots,X_n|\theta), \quad \forall \theta \in \Theta.$$

• The constant does not depend on θ :

$$\pi(\theta|X_1,\ldots,X_n) = \frac{\pi(\theta)p_n(X_1,\ldots,X_n|\theta)}{\int_{\Theta} p_n(X_1,\ldots,X_n|t) \,\mathrm{d}\pi(t)}, \quad \forall \theta \in \Theta.$$

The Bayes rule and the posterior distribution (3)

In the previous example:

•
$$\pi(p) \propto p^{a-1}(1-p)^{a-1}, p \in (0,1).$$

► Given $p, X_1, \dots, X_n \stackrel{i.i.d.}{\sim} Ber(p)$, so $p_n(X_1, \dots, X_n | \theta) = p^{\sum_{i=1}^n X_i} (1-p)^{n-\sum_{i=1}^n X_i}.$

Hence,

$$\pi(\theta|X_1,\ldots,X_n) \propto p^{a-1+\sum_{i=1}^n X_i} (1-p)^{a-1+n-\sum_{i=1}^n X_i}$$

The posterior distribution is

$$\mathcal{B} \quad a + \sum_{i=1}^{n} X_i, a + n - \sum_{i=1}^{n} X_i$$

.

Non informative priors (1)

- Idea: In case of ignorance, or of lack of prior information, one may want to use a prior that is as little informative as possible.
- Good candidate: $\pi(\theta) \propto 1$, i.e., constant pdf on Θ .
- If Θ is bounded, this is the uniform prior on Θ .
- If Θ is unbounded, this does not define a proper pdf on Θ !
- An *improper prior* on Θ is a measurable, nonnegative function $\pi(\cdot)$ defined on Θ that is not integrable.
- In general, one can still define a posterior distribution using an improper prior, using Bayes' formula.

Non informative priors (2)

Examples:

• If
$$p \sim \mathcal{U}(0,1)$$
 and given $p, X_1, \ldots, X_n \stackrel{i.i.d.}{\sim} Ber(p)$:
 $\pi(p|X_1, \ldots, X_n) \propto p^{\sum_{i=1}^n X_i} (1-p)^{n-\sum_{i=1}^n X_i},$

i.e., the posterior distribution is

$$\mathcal{B} \quad 1 + \sum_{i=1}^{n} X_i, 1 + n - \sum_{i=1}^{n} X_i$$
.

• If $\pi(\theta) = 1, \forall \theta \in \mathbb{R}$ and given $\theta, X_1, \dots, X_n \stackrel{i.i.d.}{\sim} \mathcal{N}(\theta, 1)$:

$$\pi(\theta|X_1,\ldots,X_n) \propto \exp \left(-\frac{1}{2} \prod_{i=1}^n (X_i - \theta)^2\right),$$

i.e., the posterior distribution is

$$\mathcal{N}\left(\bar{X}_n, \frac{1}{n}\right).$$

Non informative priors (3)

Jeffreys prior:

$$\pi_J(\theta) \propto \sqrt{\det I(\theta)},$$

where $I(\theta)$ is the Fisher information matrix of the statistical model associated with X_1, \ldots, X_n in the frequentist approach (provided it exists).

- In the previous examples:
 - Ex. 1: $\pi_J(p) \propto \frac{1}{\sqrt{p(1-p)}}$, $p \in (0,1)$: the prior is $\mathcal{B}(1/2, 1/2)$.

• Ex. 2: $\pi_J(\theta) \propto 1$, $\theta \in {\rm I\!R}$ is an improper prior.

Non informative priors (4)

▶ Jeffreys prior satisfies a reparametrization invariance principle: If η is a reparametrization of θ (i.e., η = φ(θ) for some one-to-one map φ), then the pdf π̃(·) of η satisfies:

$$\tilde{\pi}(\eta) \propto \sqrt{\det \tilde{I}(\eta)}$$

where $\tilde{I}(\eta)$ is the Fisher information of the statistical model parametrized by η instead of θ .

Bayesian confidence regions

▶ For $\alpha \in (0,1)$, a Bayesian confidence region with level α is a random subset \mathcal{R} of the parameter space Θ , which depends on the sample X_1, \ldots, X_n , such that:

$$\mathbb{P}[\theta \in \mathcal{R} | X_1, \dots, X_n] = 1 - \alpha.$$

- Note that \mathcal{R} depends on the prior $\pi(\cdot)$.
- "Bayesian confidence region" and "confidence interval" are two distinct notions.

Bayesian estimation (1)

- The Bayesian framework can also be used to estimate the true underlying parameter (hence, in a frequentist approach).
- In this case, the prior distribution does not reflect a prior belief: It is just an artificial tool used in order to define a new class of estimators.
- ▶ Back to the frequentist approach: The sample X_1, \ldots, X_n is associated with a statistical model $(E, (\mathbb{P}_{\theta})_{\theta \in \Theta}).$
- Define a distribution (that can be improper) with pdf π on the parameter space Θ.
- Compute the posterior pdf π(·|X₁,...,X_n) associated with π, seen as a prior distribution.

Bayesian estimation (2)

Bayes estimator:

$$\hat{\theta}^{(\pi)} = \int_{\Theta} \theta \,\mathrm{d}\pi(\theta | X_1, \dots, X_n) :$$

This is the *posterior mean*.

The Bayesian estimator depends on the choice of the prior distribution π (hence the superscript π).

Bayesian estimation (3)

- In the previous examples:
 - Ex. 1 with prior $\mathcal{B}(a, a)$ (a > 0):

$$\hat{p}^{(\pi)} = \frac{a + \sum_{i=1}^{n} X_i}{2a + n} = \frac{a/n + \bar{X}_n}{2a/n + 1}.$$

In particular, for a = 1/2 (Jeffreys prior),

$$\hat{p}^{(\pi_J)} = \frac{1/(2n) + \bar{X}_n}{1/n + 1}.$$

• Ex. 2: $\hat{\theta}^{(\pi_J)} = \bar{X}_n$.

- In each of these examples, the Bayes estimator is consistent and asymptotically normal.
- In general, the asymptotic properties of the Bayes estimator do not depend on the choice of the prior.

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